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Lampiran 1. Data Penelitian Jumlah Kasus Mortalita Pasien Difteri di Indonesia Tahun 2019

No	Provinsi	Y	X1	X2	X3	X4	X5	X6
1	Aceh	0	15.32	6.3	70.7	10.08	69.76	75
2	Sumatera Utara	3	8.83	5.8	73.86	10.24	58.66	199
3	Sumatera Barat	0	6.42	3.9	73.09	37	71.59	81
4	Riau	0	7.08	4.4	69.56	43.24	62	78
5	Jambi	0	7.6	4.9	62.94	40.25	67.18	40
6	Sumatera Selatan	0	12.71	3.9	63.26	37.94	69.68	73
7	Bengkulu	0	15.23	2.4	49.52	35.05	66.63	29
8	Lampung	0	12.62	2.7	59.65	31.6	70.58	99
9	Kep. Bangka Belitung	1	4.62	4.6	84.09	59.37	80.21	23
10	Kepulauan Riau	0	5.9	2.7	91.53	37.73	57.2	29
11	DKI Jakarta	1	3.47	1.5	88.44	61.25	64.55	191
12	Jawa Barat	10	6.91	2.5	72.38	28.84	65.44	330
13	Jawa Tengah	0	10.8	3.5	78.86	44.69	83.15	295
14	D I Yogyakarta	0	11.7	1.4	75.3	66.21	71.95	75
15	Jawa Timur	1	10.37	3.6	76.1	44.5	37.18	372
16	Banten	2	5.09	3	73.68	24.79	46.88	121
:		:	:	:	:	:	:	:
:		:	:	:	:	:	:	:
21	Kalimantan Tengah	0	4.98	4.7	69.18	42.97	67.27	30
22	Kalimantan Selatan	0	4.55	3.9	72.05	34.64	69.07	44
23	Kalimantan Timur	0	5.94	3.4	93.77	47.79	43.46	53
24	Kalimantan Utara	0	6.63	1.7	80.58	57.47	73.17	8
25	Sulawesi Utara	0	7.66	6.3	81.72	17.2	19.47	47
26	Sulawesi Tengah	0	13.48	3.8	74.67	50.84	58.41	38
27	Sulawesi Selatan	0	8.69	4.8	79.32	36.63	45.26	111
28	Sulawesi Tenggara	2	11.24	6.3	79.32	39.39	39.68	36
29	Gorontalo	0	15.52	8.1	82.33	26.36	45.72	14
30	Sulawesi Barat	0	11.02	5.7	63.55	38.32	45.28	12
31	Maluku	0	17.69	10.3	73.17	45.01	48.15	30
32	Maluku Utara	0	6.77	6.5	64.82	44.67	78.7	21
33	Papua Barat	1	22.17	4.1	70.18	42.55	55.97	19
34	Papua	0	27.53	4.5	39.05	20	34.12	44

Lampiran 2. Turunan Pertama dan kedua

- Fungsi Z_i

$$Z_i = Z(\lambda_i, \phi) = \sum_{j=0}^{\infty} \frac{\exp(\mathbf{x}_i^T \boldsymbol{\beta})^j}{(j!)^\phi}$$

Turun pertama Fungsi Z_i terhadap ϕ

$$\begin{aligned} \frac{\partial Z_i}{\partial \phi} &= \frac{\partial Z_i}{\partial \phi} \left(\sum_{j=0}^{\infty} \frac{e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j}}{(j!)^\phi} \right) \\ &= \frac{\partial Z_i}{\partial \phi} \left(\sum_{j=0}^{\infty} e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j} (j!)^{-\phi} \right) \\ &= \sum_{j=0}^{\infty} e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j} \left(\frac{\partial Z_i}{\partial \phi} \left(\sum_{j=0}^{\infty} (j!)^{-\phi} \right) \right) \\ &= \sum_{j=0}^{\infty} e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j} \left(\frac{\partial Z_i}{\partial \phi} e^{(-\phi) \ln(j!)} \right) \\ &= \sum_{j=0}^{\infty} e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j} \left(e^{(-\phi) \ln(j!)} \left(\frac{\partial Z_i}{\partial \phi} ((-\phi) \ln(j!)) \right) \right) \\ &= \sum_{j=0}^{\infty} e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j} \left(e^{(-\phi) \ln(j!)} \left(\frac{\partial Z_i}{\partial \phi} ((-\phi) \ln(j!)) \right) \right) \\ &= \sum_{j=0}^{\infty} e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j} (j!)^{-\phi} (\ln(j!)) \\ &= \sum_{j=0}^{\infty} \frac{-\exp(\mathbf{x}_i^T \boldsymbol{\beta})^j \ln(j!)}{(j!)^\phi} \end{aligned}$$

Turunan Kedua Fungsi Z_i terhadap ϕ

$$\frac{\partial^2 Z_i}{\partial \phi^2} = \frac{\partial Z_i}{\partial \phi} \left(\sum_{j=0}^{\infty} \frac{-\exp(\mathbf{x}_i^T \boldsymbol{\beta})^j \ln(j!)}{(j!)^\phi} \right)$$

$$\begin{aligned}
&= \sum_{j=0}^{\infty} -e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j \ln(j!)} \left(\frac{\partial Z_i}{\partial \phi} (j!)^{-\phi} \right) \\
&= \sum_{j=0}^{\infty} -e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j \ln(j!)} \left(\frac{\partial Z_i}{\partial \phi} (e^{(-\phi) \ln(j!)}) \right) \\
&= \sum_{j=0}^{\infty} -e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j \ln(j!)} (e^{(-\phi) \ln(j!)}) \left(\frac{\partial Z_i}{\partial \phi} ((-\phi) \ln(j!)) \right) \\
&= \sum_{j=0}^{\infty} -e^{(\mathbf{x}_i^T \boldsymbol{\beta})^j \ln(j!)} (j!)^{-\phi} (-\ln(j!)) \\
&= \sum_{j=2}^{\infty} \frac{\exp(\mathbf{x}_i^T \boldsymbol{\beta})^j \ln(j!)^2}{(j!)^{\phi}}
\end{aligned}$$

Lampiran 3A. Pengujian Distribusi Poisson

One-Sample Kolmogorov-Smirnov Test

		Y
N		34
Poisson Parameter ^{a,b}	Mean	.6765
Most Extreme Differences	Absolute	.227
	Positive	.227
	Negative	-.029
Kolmogorov-Smirnov Z		1.323
Asymp. Sig. (2-tailed)		.060

a. Test distribution is Poisson.

b. Calculated from data.

Lampiran 3B. Output SAS Overdispersi

Model Information

Data Set	WORK.DIFTERI
Distribution	Poisson
Link Function	Log
Dependent Variable	Y
Observations Used	34

Criteria For Assessing Goodness Of Fit

Criterion	DF	Value	Value/DF
Deviance	27	33.9971	1.2592
Scaled Deviance	27	27.0000	1.0000
Pearson Chi-Square	27	126.5310	4.6863
Scaled Pearson X2	27	100.4892	3.7218
Log Likelihood		-7.5590	
Full Log Likelihood		-28.4936	
AIC (smaller is better)		70.9871	
AICC (smaller is better)		75.2948	
BIC (smaller is better)		81.6717	

Lampiran 4. Multikolinieritas

Coefficients ^a								
Model		Unstandardized Coefficients		Standardized Coefficients	t	Sig.	Collinearity Statistics	
		B	Std. Error	Beta			Tolerance	VIF
1	(Constant)	2.536	4.107		.618	.542		
	X1	-.059	.063	-.184	-.940	.356	.633	1.581
	X2	-.049	.186	-.052	-.261	.796	.607	1.647
	X3	-.023	.028	-.161	-.831	.414	.645	1.551
	X4	-.018	.027	-.135	-.677	.505	.612	1.633
	X5	-.006	.023	-.051	-.268	.791	.678	1.476
	X6	.009	.004	.468	2.585	.016	.740	1.351

a. Dependent Variable: Y

Lampiran 5. Hasil Estimasi Parameter Model Regresi NB Menggunakan *Software SAS*

```
Data Zicmp;
Input y X1 X2 X3 X4 X5 X6;
Datalines;
Run;
proc nlmixed data= Zicmp tech= newrap;
parameters b0=0 b1=0 b2=0 b3=0 b4=0 b5=0 b6=0 nu=0;
bpart=b0+b1*X1+b2*X2+b3*X3+b4*X4+b5*X5+b6*X6;
lambda=exp(bpart);
Z = 1;
DO s = 1 to 100;
Q = 1;
DO r = 1 to s;
Q = Q * (lambda / (r**nu));
END;
Z = Z + Q;
END;
ll = (Y*log(lambda)) - (nu*lgamma(Y+1)) - (log(Z));
model Y ~ general(ll);
predict _ll out=LL_1;
run;
```

Lampiran Hasil Estimasi Parameter Model Regresi NB Menggunakan Software SAS (lanjutan)

The NLMIXED Procedure									
Specifications									
Data Set		WORK.ZICMP							
Dependent Variable		Y							
Distribution for Dependent Variable		General							
Optimization Technique		Newton-Raphson							
Fit Statistics									
-2 Log Likelihood		59.7							
AIC (smaller is better)		87.6							
AICC (smaller is better)		114.5							
BIC (smaller is better)		105.9							
Parameter Estimates									
Parameter	Estimate	Standard Error	D F	t Value	Pr > t	95% Confidence Limits	Gradient		
Intercept	2.3431	2.7724	34	0.85	0.4039	- 7.9774 3.2911	3.697E-6		
beta1	-0.1358	0.08162	34	- 1.66	0.1053	- 0.0300 0.3017	0.0000285		
beta2	0.04772	0.1848	34	0.26	0.7978	- 0.4233 0.3279	0.000011		
beta3	- 0.02994	0.02136	34	- 1.40	0.1702	- 0.0134 0.0733	0.0002686		
beta4	- 0.02229	0.01795	34	- 1.24	0.2228	- 0.0141 0.0587	0.0001117		
beta5	- 0.00209	0.01515	34	- 0.14	0.8912	- 0.0287 0.0328	0.0002288		

b6	0.00604 3	0.002539	34	2.38	0.0231	0.0008 82	0.0112 0	0.001174
Nu	0.2916	0.3531	34	0.83	0.4145	- 0.4259	1.0091	-9.66E-6

Lampiran 6. Sintaks MATLAB untuk Estimasi Parameter

```
%% Ambil dataset
dataset = xlsread('F:\DIFTERI1.xlsx');
[n0,p0] = size(dataset);
datafull = dataset(:,1:6);
%% Definisi X dan Y
Y = datafull(:,1);
X = datafull(:,2:6);
[n,p] = size(X);
%% Matrix X baru
Xbaru = [ones(n,1),X];
X1 = Xbaru(:,1);
[nb,kb] = size(Xbaru);
%% V, Beta dan Gamma
B= (Xbaru'*Xbaru)\(Xbaru'*Y);
GM = (Xbaru'*Xbaru)\(Xbaru'*Y);
T0 = [v,B',GM'];
%% iterasi & selisi
iterasi1 = 0;
iterasi2 = 0;
selisih1 = 0.1;
selisih2 = 0.1;

while selisih1 >= 0.0001
while selisih2 >= 0.0001
    XB = Xbaru*B;
    XG = Xbaru*GM;
    LM= exp(XB);

    %% E-Step
    for j=1:kb
        for i= 1:n
            Zi(i)= LM(i)^i/(factorial(i))^v;
            % Turunan pertama
            dldv(j)=-log(factorial(Y(j)))-(dZdv(j)/Zi(i));
            % Turunan pertama B
            dZdB(i,j)= j*Xbaru(i,j)*LM(i).^j/(factorial (j))^v;
            dldB(i,j)= Y(i)*Xbaru(i,j)-(dZdB(i,j))/Zi(i);
        end
        for ii= 1:n
            L(ii) = XG(ii)+LM(ii);
            z(ii) = exp(XG(ii))/(1+exp(-L(ii)));
            dBB(ii,j) = (z(ii)-1)*dldB(ii,j);
            dVV(j)= (z(ii)-1)*dldv(j);
            dGG(ii,j)= z(ii)*Xbaru(ii,j);
        end
        for iii=1:n
            dG(iii,j) = -Xbaru(iii,j)*exp(XG(iii))/(1+exp(XG(iii)));
        end
        d1v= sum(dVV);
        d1B(j)= sum(dBB(:,j));
        d1G(j)= sum(dGG(:,j)) + sum(dG(:,j));
    end
end
```

```

end
for k=1:kb
    for i=1:kb
        for jj=2:kb
            Zi(i) = (LM(i)^i)/(factorial(i))^v;
            % Turunan kedua v^2
            d2ldv(k) = (((dZdv(k))^2)/(Zi(i)^2))-((d2Zdv(k))/Zi(i));
            % Turunan pertama B^2
            d2ldb(i,k) = ((dZdB(i,k)^2)/Zi(i))-d2ZdB(k)/Zi(i);
            % turunan kedua v dan B
            d2VB(i,k) = ((dZdv(k)*dZdB(i,k))/(Zi(i)^2))-(d2ZvB(i,k)/Zi(i));
        end
        for ii=1:n
            L(ii)=XG(ii)+LM(ii);
            z(ii) = exp(XG(ii))/(1+exp(-L(ii)));
            d2BB(ii,k) = (z(ii)-1)*d2ldb(i,k);
            d2VV(ii,k)=(z(ii)-1)*d2ldv(k);
            d2VBB(ii,k)=(z(ii)-1)*d2VB(k);
        end
        for iii= 1:n
            EG(iii) = exp(XG(iii))/(1+exp(XG(iii)));
            d2G(iii,j,k) = Xbaru(iii,j)*Xbaru(iii,k)*(EG(iii)^2- EG(iii));
        end
        d2LV= sum(d2ldv);
        d2LVB(j,k) = sum(d2VB(:,k));
        d2LBB(j,k) = sum(d2BB(:,k));
        d2LGG(j,k) = sum(d2G(:,j,k));
    end
end

T = [v,B'];
G = [d1v;d1B'];
H1= [d2LV,d2LVB(6,:)];
H2= [d2LVB(6,:)',d2ldb];
H = [H1;H2];

%% M-step
Tlama= T';
iterasi1=iterasi1+1;
c =inv(H)*G;
Theta= Tlama-c;
Selisih1 =norm(Theta-Tlama);
disp([iterasi1' Theta' Selisih1'])
Glama = GM;
iterasi2 = iterasi2 + 1;
d =inv(d2LGG)*d1G';
Gamma = Glama - d;
Selisih2 = norm(Gamma - Glama);
disp([iterasi2' Gamma' Selisih2'])

LM = Theta(1,:);
B = Theta(1,:);
GM = Gamma;
end

```

```

end

%% Uji Parsial Parameter Model
VB= -(inv(H));
VGM = -(inv(d2LGG));
% Uji Parsial Parameter Beta
for w=1:kb
    SEB(w) = sqrt(VB(j,j));
    WB(w) = (Gamma(w,:)/SEB(:,w))^2;
end
% Uji Parsial Parameter Gamma
for j=1:kb
    SEGM(j) = sqrt(VGM(j,j));
    WGM(j) = (Gamma(j,:)/SEGM(:,j))^2;
end

```